DIVERSIFICATION BENEFITS OF INCLUDING TURKISH AND US STOCKS IN A PORTFOLIO

Engin Küçükkaya METU Department of Business Administration Ankara, Turkey engink@metu.edut.r

Abstract

Portfolio diversification is an important issue in the investment process. Country diversification is a way to achieve higher returns with lower risk for international portfolio investment. In this study, I examine possible short and long term relationships between the US and the Turkish equity markets by employing correlation analysis, cointegration methodology and Granger causality. Results indicate that the two markets are not highly correlated and there is no cointegrating relationship, pointing to possible diversification benefits by investing in the Turkish market. However, long term Granger causality results that show influence of the US market on the Turkish market limits the benefits that can be derived from diversification.

Key Words: Portfolio, Portfolio Diversification, Equity Markets.

1. Introduction

In the context of "globalization", world economies are becoming more interconnected and integrated. Increasing volume of international trading, and even faster increase of the international capital flows are the two main drivers of globalization. Trade liberalization and increasing number of memberships with the WTO can be listed as the reasons for larger international trade volume, where the liberalization and deregulation in the capital markets, elimination of foreign exchange controls, and search for higher returns of capital exporting economies can be listed as the reasons for larger international capital flows.

The issue about the higher level of integration and interrelation among world capital markets, especially world stock markets, has received a great level of attention by the financial media, as well as the research community. Lately, when we read financial news, we can see that the movements and expectations in the stock markets are usually interpreted by including the effects of the foreign stock markets as well as

the currency markets. This is especially true for the relatively smaller markets, where the movements in the larger markets are described as the major source of influence. Turkish stock market is one of those smaller markets, and its movements are generally described in the financial media as the results of what happens in the global stock and currency markets. It is also reported that a large part (around 70%) of the Turkish stock market is controlled by foreign investors and a large part of the trading volume (e.g. 45% in the first 6 months of 2007) is realized by those. In light of these facts, it is not unexpected that the foreign investors providing a link between the world markets and the Turkish stock market. From the financial point of view, one can explain the large capital flows and ownership of foreign investors in especially developing markets such as Turkey with two motivational factors: higher returns and lower overall portfolio risk. It is an empirical issue to determine if the developing markets provide higher returns than the developed markets. In some sample periods this can be true, but in the long run, because of the higher volatility, it may not be the case, especially when considered in a common currency. On the other hand, although those developing market returns seem to have higher volatility, if their returns are not correlated with those of the developed markets, they can provide significant portfolio diversification benefits. In the past, these benefits were tied to factors such as barriers to international capital flows, exchange controls, lack of free trade, inadequate information on foreign securities and investor bias against foreign securities. However, in later years, significant increase in correlations and volatility transmission between equity markets during, and after, the 1987 equity market crash are considered as the reason for lower available benefits by international diversification.

In this study, I examine the long and short run dynamics between the world's largest stock market, the US market and the Turkish stock market in order to determine if investing in the Turkish market provides diversification benefits relative to the US market. Results indicate that the Turkish stock market is not strongly correlated to the US market in the short term, and it is not integrated to the US market in the long run, therefore it is a good choice to provide diversification. However, a long run influence of the US market is also discovered, and should be considered as a limiting influence on the benefits expected.

The rest of the paper proceeds as follows: next section provides a brief review of the related literature, followed by the data and methodology employed. Results of the empirical tests are presented in the fourth section, and the last section summarizes results and concludes.

2. Literature review of the interdependencies among financial markets

There are a large number of studies that investigate the dynamic linkages among world equity markets. Some of them provide evidence for integration, and some of them, against. Those studies that examine multiple countries provides supporting evidence for some countries and against some others.

Hamao and Masulis (1990), Kasa (1992), and Arshanapalli and Doukas (1993) present evidence for the equity markets of the developed countries being integrated. Chen et al., (2002) for emerging markets, Gilmore and McMannus (2002) between the US and three developing Central European Markets and Manning (2002) for South East Asia, all find cointegrating relationships among equity markets. In addition to those, Tokic (2003) supports the integration results for the USA and five markets; Australia, Japan, Hong Kong, New Zealand and Singapore; Cha and Oh (2000) for the Japanese, the US, Hong Kong, Korea, Singapore and Taiwan; and Ghosh et al. (1999) for the US, Hong Kong, India, Korea and Malaysia. Eun and Shim (1989) finds effects of the US equity markets on world markets, and also argue that the US is the most dominant market in the world, an argument usually supported in other studies that include the US in their sample.

Some of the studies that do not support the integrating relationship are as follows: Yang, Khan and Pointer (2003) for developed markets, DeFusco et al. 1996 for the US market and thirteen emerging capital markets, Pan et al. (1999) for six country's equity indices. No long-term linkages between the US and European equity markets were found by Byers and Peel (1993) and Kanas (1998). Ghosh et al. (1999) does not find any effect of the USA and Japan on the stock markets of Taiwan and Thailand.

A very recent study by Ozdemir et al. (2008) find long term relationship (cointegration) for 8 of the 15 emerging markets they examine. For the remaining 7, including Turkey, they do not find evidence of cointegration, but they document granger causality from the US, but not vice versa. Another study on Turkey was conducted by Berument and Ince 2005, who assess the effect of S&P500 return on the Istanbul Stock Exchange by using daily data. Their estimates suggest that returns on S&P500 affect ISE return positively up to four days.

3. Data and methodology

Data used in this study is the monthly MSCI Indices for the US and for Turkey, both denominated in US Dollars, between May 1988 and May 2008 (a total of 241 monthly observations). By using indices with a common denominator, and that are constructed with same methodology (MSCI), I make sure that these indices are comparable, and lead to meaningful results. As the indices represent the level of the equity markets for each country, I take the natural logarithm of both in order to better fulfill distributional assumptions.

The first step in the analysis is finding the average monthly returns, their standard deviations and the Sharpe Index for the two markets examined. Those measures are helpful in comparing the returns of the two markets. As a part of the analysis, they are followed by the correlation among the two markets, an indicator of short term co-movements.

In the second part, cointegration analysis is employed, in order to examine the long term co-movements. Cointegration is based on the idea that if two or more non-stationary economic variables are in an equilibrium relationship, then their stochastic trends must be linked. In other words, cointegration represents a stationary linear combination of variables that are non-stationary. That linear combination is generally considered as representing an equilibrium relationship among the factors. As a complementary measure to correlation analysis, which is meaningful in the short term, cointegration is a measure of long-term co-movement in the long term.

When two or more series are cointegrated, error correction models (ECMs) can be used to examine the short-term relationships, as stated by Johansen [1988]. Error correction models are based on the idea that cointegrated series can only have short-term deviations from equilibrium, and these deviations will be corrected in the long-run, thus the series will drift together. In ECMs, there is an error correction term for each cointegrating vector included in each equation in addition to the lagged values of variables.

The first step in cointegration analysis is to test for the presence of a unit root in each series in the system. The Augmented Dickey and Fuller [1979] (ADF) methodology is generally used for unit root tests.

However, when a series is shown to have a unit root, another round of unit root tests have to be applied to the first differences of the series in order to conclude that the series is not integrated with an order higher than one. Series with higher levels of integration should not be included in the cointegration analysis when the other series are I(1).

If the series examined are shown to have unit roots, the next step is testing for the presence of cointegration among the series. Johansen's [1988] maximum likelihood method is a widely used method used in estimating the systems.

When using Johansen's method, the system is represented by the following equation:

$$\Delta X_{t} = \Pi X_{t-p} + \sum_{i=1}^{p-l} \Gamma_{i} \Delta X_{t-i} + \mu + \varepsilon_{t}$$

where
$$\Gamma_{i} = -(I - A_{1} - ... - A_{i})$$

and
$$\Pi = -(I - A_1 - ... - A_p)$$
.

 X_t is the vector of the variables, and Δ is the first difference operator. p is the maximum number of lags, and ε_t is a four dimensional identically independently distributed vector with zero mean and variance matrix Ω_{ε} . The n×n Π matrix, the long run impact matrix, can be represented as the multiplication of two n×r matrices α and β as in $\Pi = \alpha\beta'$. r represents the rank of Π matrix, and it is less than or equal to n.

The test for cointegration is a rank test for Π matrix. Reduced rank of Π implies that under certain conditions the process ΔX_t is stationary, X_t is nonstationary, but also that $\beta'X_t$ is stationary. Thus we can interpret the relations $\beta'X_t$ as the stationary relations among nonstationary variables, i.e. as cointegrating relations.

The rank test for Π is the test for the number of characteristic roots of the Π matrix that are significantly different from zero, and it is conducted using the following log-likelihood ratio test statistics.

The first test statistic, λ_{max} , tests the null that the number of cointegrating vectors is r against the alternative of r+1 cointegrating vectors. The second statistic, λ_{trace} , tests the null hypothesis that the number of cointegrating vectors is less than or equal to r against a general alternative. λ_{max} is preferred as the primary selection criteria to pin down the number of cointegrating vectors, as it provides a sharp alternative

hypothesis. However, λ_{trace} statistic is conventionally reported and used as a supporting test for choosing the number of cointegrating vector, as it may provides a more robust test statistic in some cases. The critical values for the two tests are provided in Johansen and Juselius [1990].

When the cointegration tests result show no cointegrating among variables, we can use the VAR system in first differences to examine short-term influences. However, differencing the series lead to loss of information that is contained in the levels. In order not to lose any information that is present in the levels, I employ an alternative methodology, the Toda-Yamamoto (1995) procedure to test for long run Granger causality. This procedure employs a VAR system in levels, consequently it does not lead to a loss of information due to differencing. On the negative side, it has limited results as the procedure allows for long run Granger causality tests only.

4. Empirical Results

Monthly average stock index returns, standard deviations of the returns, and Sharpe index values (using US 3 month T-Bill rate as risk free rate) of the US and Turkish equity market indices are presented in *Table 1* below.

Table 1. Mean return, sd, and Sharpe index

	US	TR
Mean Return	0,72%	0,82%
Standard Deviation	0,0396	0,1774
Sharpe	0,0914	0,0259
Correlation	0.2252	

The mean monthly return in Turkish market is higher than the US market, and the standard deviation of the Turkish market returns is also much higher than the standard deviation of the US market returns. These numbers indicate that Turkish market may provide higher average return, albeit accompanied by a considerable level of risk. A better comparison can be based on the Sharpe index, which measures the extra return provided for a unit (total) risk. When we examine the Sharpe index values, it is clear that the extra return provided by the Turkish market is not enough for bearing the extra risk, when compared to the US market.

When the short term co-movements between the two markets are considered, we can see that the two markets are not highly correlated, with a correlation coefficient of less than 23%. This value is an indication that having these two markets in a portfolio can provide diversification benefits in the short term.

As the next step, I examine the possibility of long-term diversification benefits, by employing cointegration analysis. Before proceeding to the next level, I first tested the series for unit roots by using two popular tests: augmented Dickey and Fuller (1979) (ADF), Phillips and Perron (1988) (PP). The unit root test results are shown in Table 2.

Table 2. Unit root test results

	LEVELS				
		ADF	PP		
	LTU	-2.022814(0)	-2.189575		
Intercept	LUS	-1.716258(0)	-1.716258		
Intercept	LTU	-2.746078(0)	-3.037539		
and Trend	LUS	-1.247228(0)	-1.238562		
	FIRST DIFFERENCES				
		ADF	PP		
	LTU	-14.71834 ^a (0)	-14.71655 ^a		
Intercept	LUS	-15.60480 ^a (0)	-15.60467 ^a		
Intercept	LTU	-14.68701 ^a (0)	-14.68510 ^a		
and Trend	LUS	-15.66517 ^a (0)	-15.66517 ^a		
denotes significance at 1% level					

LTR is the log of MSCI Turkey series (in US\$)

LUS id the log of MSCI US series (in US\$)

Both unit root tests yield similar results, that shows that LTR and LUS are both I(1). Hence, the maximum order of integration (d) is determined to be 1.

As both variables are integrated of the same order, Johansen cointegration tests are run by employing the two series.

Table 3. Cointegration test results

		Number of cointegrating vectors	
		0	<= 1
intercept and no			
deterministic trend	λ_{max}	7.474093	2.624913
	λ_{trace}	10.09901	2.624913
intercept and trend in CE	λ_{max}	8.880922	2.676565
no trend in VAR	λ_{trace}	11.55749	2.676565
intercept and trend in both	λ_{max}	8.863518	0.880856
	λ_{trace}	9.744374	0.880856

Based on the results presented in Table 3, there is no cointegration between the Turkish and US stock market series for the period under examination. These results lead to the conclusion that the Turkish market, namely IMKB is a good candidate for portfolio diversification for the long term investors that holds portfolios from the US market, confirming the result obtained for the short term.

As there was no evidence of cointegration, I couldn't run the VEC Models to examine the Granger causality. Instead, I employ the Toda-Yamamoto procedure, where the long-run Granger causality test results are shown in Table 4.

Table 4. Granger non-causality test results

Hypothesis	Wald statistic
	4.806736 ^b
LUS does not Granger cause LTU	(0.0293)
-	0.908042
LTU does not Granger cause LUS	(0.3416)

The p-values are in parentheses. Superscript b stands for significance at 5%

These results indicate that the US stock market Granger causes Turkish stock market in the long run, somewhat limiting the benefits that can be derived from having both markets in an investment portfolio.

5. Conclusion

In this study I examine the short run and the long run relationship between the American and Turkish equity markets. The results show that the two markets do not have high level of correlation in the short run, and also they are not in a long term equilibrium relationship, indicated by the lack of cointegration. Implication is that the Turkish market can be used to achieve diversification benefits when used with a portfolio of investment that includes the US market, both in the short term and the long term. However, as indicated by the long run Granger causality results, the US market has an influence on the Turkish market, limiting the benefits that can be achieved in terms of diversification.

References

- Arshanapalli, B., Doukas, J. 1993, International stock market linkages: evidence from the pre- and post-October 1987 period, Journal of Banking and Finance 17: 193–208.
- Berument H., Ince O. 2005, Effect of S&P500 return on emerging markets:Turkish Experience, Applied Financial Economics Letters 1:1: 59–64
- Byers, M. D. and Peel, D.A.1993, Some evidence on the interdependence of national stock markets and the gains from international portfolio diversification, Applied Financial Economics 3: 239–42.
- Cha, B., Oh, S. 2000, The relationship between developed equity markets and the Pacific Basin's emerging equity markets, International Review of Economics and Finance 9: 299–322.
- Chen, G., Firth, M., Rui, O.M. 2002, Stock market linkages: evidence from latin America, Jpornal of Banking and Finance 26: 1113–1141.
- De Fusco, R.A., Geppert, J.M., Tsetsekos, G.P. 1996, Long-run diversification potential in emerging stock markets, The Financial Review 31: 343–363.

- Dickey, D. A., and Fuller, W. A. 1979, Distribution of the Estimators for Autoregressive Time Series with a Unit Root, Journal of the American Statistical Society 75: 427-431.
- Eun, C. S. and Shim, S. 1989, International transmission of stock market movements, Journal of Financial and Quantitative Analysis 24: 241–56.
- Gilmore, C.G., McMannus, G.M. 2002, International portfolio diversification: US and central European equity markets, Emerging Markets Review 3: 69–83.
- Ghosh, A., Saidi, R. and Johnson, K. H. 1999, Who moves the Asia-Pacific stock markets US or Japan? Empirical evidence based on the theory of co-integration, The Financial Review 34: 159–70.
- Hamao, Y.R., Masulis, V.Ng. 1990, Correlation in price changes and volatility across international stock markets, Review of Financial Studies 3: 281–307.
- Johansen, S. 1991, Estimation and hypothesis testing of cointegration vectors in Gaussian vector autoregressive models, Econometrica, **59**: 1551–1580.
- Johansen, S. 1992, A determination of the co-integration rank in the presence of a linear trend, Oxford Bulletin of Economics and Statistics, 54: 383–397.
- Johansen, S., and Juselius, K. 1990, Maximum Likelihood Estimation and Inference on Cointegration-With Applications to the Demand for Money, Oxford Bulletin of Economics and Statistics, 52: 169-210.
- Kanas, A. 1998, Linkages between the US and European equity markets: further evidence from cointegration tests, Applied Financial Economics 8: 607–14.
- Kasa, K. 1992, Common stochastic trends in international stock markets, Journal of Monetary Economics 29: 95–124.
- Manning, N. 2002, Common trends and convergence? South East Asian equity markets, 1988-1999, Journal of International Money and Finance 21: 183-202
- Ozdemir, Z.A., Olgun, H., Saracoglu, B. 2008, Dynamic linkages between the center and periphery in international stock markets, Research in International Business and Finance: (articles in press), doi:10.1016/j.ribaf.2008.05.001

- Pan, M-S., Liu, Y.A., Roth, H.J. 1999: Common stochastic trends and volatility in Asian–Pacific equity markets, Global Finance Journal 10: 161–172.
- Phillips, P.C.B., Perron, P. 1988, Testing for a unit root in time series regression, Biometrika 75: 335-346.
- Toda HY, Yamamoto T. 1995, Statistical inference in vector autoregression with possibly integrated processes, Journal of Econometrics 66: 225-250.
- Tokic, D. 2003, Emerging markets before the 1997 Asia Pacific financial crisis, Asia Pacific Business Review 9(3): 105–15.
- Yang J., Khan M.M., Pointer L. 2003, Increasing Integration Between the United States and Other International Stock Markets?, Emerging Markets Finance and Trade 39(6): 39–53.