# GENERALIZED ASYMMETRIC POWER ARCH MODELING OF NATIONAL STOCK MARKET RETURNS

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#### Özet

Uygulamalı çalışmalar finansal varlık getirilerinin şişman kuyruk (leptokurtosis) özelliği sergilediklerini ve genellikle oynaklık kümelenmesi ve asimetrik yapı ile nitelendirildiklerini göstermiştir. Bu çalışmada, sekiz ülkenin ulusal borsa endeks getirilerinde (Nasdaq100, DAX, Nikkei225, Strait Times, MerVal, IPC, Shanghai Composite and ISE100) farklı hata dağılımlarına bağlı olarak oynaklık yapılarını belirlemek üzere Ding, Granger and Engle (1993) tarafından ileri sürülen Genelleştirilmiş Asimetrik Üslü ARCH (APGARCH) modelinin uygulanabilirliği araştırılmıştır. Çalışmanın bulguları, piyasalarda yaşanan gelişmelere karşı asimetrik etkilerin varlığında, finansal zaman serilerindeki çarpıklık ve basıklık özelliklerini birlikte ele alan çarpık Student-*t* dağılımlı APGARCH(1,1) modelinin tercih edilmesi gerektiği yönündedir.

Anahtar Kelimeler: APGARCH, çarpık Student-t dağılımı, borsa endeks getirileri

JEL Sınıflandırması: C22, C52, G12, G15

#### Abstract

Empirical studies have shown that a large number of financial asset returns exhibit fat tails (leptokurtosis) and are often characterized by volatility clustering and asymmetry. This paper considers the ability of the Generalized Asymmetric Power ARCH (APGARCH) model introduced by Ding, Granger and Engle (1993) to capture the stylized features of volatility in national stock market returns for eight countries (Nasdaq100, DAX, Nikkei225, Strait Times, MerVal, IPC, Shanghai Composite ve ISE100). The results of this paper suggest that in the presence of asymmetric responses to innovations in the market, the APGARCH(1,1) Skewed Student-*t* model which accommodates both the skewness and the kurtosis of financial time series is preferred.

Keywords: APGARCH, skewed Student-t distribution, stock market returns

JEL Code: C22, C52, G12, G15

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## INTRODUCTION

Asset returns are approximately uncorrelated but not independent through time as large (small) price changes tend to follow large (small) price changes. This temporal concentration of volatility is commonly referred to as volatility clustering and it was not fully exploited for modeling purposes until the introduction of the ARCH model by Engle (1982) and Generalized ARCH (GARCH) model by Bollerslev (1986). There have been numerous developments in the ARCH literature to refine both the mean and variance equations, in order to better capture the stylized features of high frequency data. A common feature of the standard class of ARCH models is that they relate the conditional variance to lagged squared residuals and past variances.

The ARCH literature has developed so rapidly. One recent development in the ARCH literature has focused on the power term by which the data is to be transformed. Ding, Granger and Engle (1993) introduced a new class of ARCH model called the Power ARCH model which estimates the optimal power term. They also found that the absolute returns and their power transformations have a highly significant long-term memory property as the returns are highly correlated.

Another important innovation has been development of ARCH model specifications to describe the asymmetry present in financial data. Stock market returns data commonly exhibits an asymmetry in that positive and negative shocks to the market do not bring forth equal responses. This phenomenon most commonly attributed to the leverage effect (see Black 1976, Christie 1982 and Nelson 1991). The applicability of the Power ARCH class of model to stock market data has been well documented in papers such as Ding, Granger and Engle (1993), Hentschel (1995), Giot and Laurent (2003) and, Pan and Zhang (2006).

Because of the empirical studies have shown that a large number of financial asset returns exhibit fat tails (leptokurtosis) and asymmetry in volatility, the main purpose of this paper is to examine the adequacy of the APGARCH model to capture the stylized features of volatility in national stock market returns for eight countries. The results of this paper suggest that in the presence of asymmetric responses to innovations in the market, the APGARCH(1,1) Skewed Student-*t* model is preferred. However, as internal dynamics of each market are different, there is inability to judge the parameters of the model by distinguishing markets as developed and emerging markets.

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The remainder of this paper proceeds as follows. In section 2 details the general model and discusses how various ARCH models are nested within this APGARCH structure. Section 3 describes the national stock market returns data to be used in this study and presents the empirical results. The robustness of these findings is assessed using the Akaike Information Criterion (AIC) and log-likelihood (LL) values. Section 4 contains some concluding remarks.

# METHODOLOGY

The Generalized Asymmetric Power ARCH (APGARCH) model, which was introduced by Ding, Granger and Engle (1993), is presented in the following framework:

$$y_t = c_0 + \varepsilon_t \tag{1}$$

$$\varepsilon_t = z_t \sigma_t \tag{2}$$

$$z_t \overset{\text{i.i.d.}}{\Box} f(0,1) \tag{3}$$

$$\sigma_t^{\delta} = \omega_0 + \sum_{i=1}^p \alpha_i (|\varepsilon_{t-i}| - \gamma_i \varepsilon_{t-i})^{\delta} + \sum_{j=1}^q \beta_j \sigma_{t-j}^{\delta}$$
(4)

where  $c_0$  is a constant parameter,  $\varepsilon_i$  is the innovation process,  $\sigma_i$  is the conditional standard deviation,  $z_i$  is an independently and identically distributed (i.i.d.) process. f(.) is the probability density function (PDF) and F(.) is the cumulative density function (CDF) with  $\omega_0 > 0, \alpha_i \ge 0, \beta_j \ge 0, \delta \ge 0$  and  $|\gamma_i| \le 1$ . Here  $\alpha_i$  and  $\beta_j$  are the standard ARCH and GARCH parameters,  $\gamma_i$  is the leverage parameter and  $\delta$  is the parameter for the power term. A positive (resp. negative) value of the  $\gamma_i$ means that past negative (resp. positive) shocks have a deeper impact on current conditional volatility than past positive (resp. negative) shocks.

The model imposes a Box and Cox (1964) transformation in the conditional standard deviation process and the asymmetric absolute innovations. In the APGARCH model, good news ( $\varepsilon_{t-i} > 0$ ) and bad news ( $\varepsilon_{t-i} < 0$ ) have different predictability for future volatility, because the conditional variance depends not only on the magnitude but also on the sign of  $\varepsilon_i$ . To put Equation (4) into operation we need to specify the lag structure and in this paper a first order lag structure is adopted for both the ARCH and GARCH terms:

$$\sigma_{t}^{\delta} = \omega + \alpha_{1} \left( \left| \varepsilon_{t-1} \right| - \gamma \varepsilon_{t-1} \right)^{\delta} + \beta_{1} \sigma_{t-1}^{\delta}$$
(5)

where  $\omega, \alpha_1, \gamma_i, \beta_1$  and  $\delta$  are additional parameters to be estimated. Equation (5) shall hereafter be referred to as a Generalized Asymmetric Power ARCH (APGARCH) model to reflect the inclusion of the  $\beta$  term. Thus, we are able to distinguish this model from a version in which  $\beta_1 = 0$ , that we shall refer to as an Asymmetric Power ARCH (APARCH) model.

In the influential paper of Engle (1982), the density function of  $z_t$ , f(.) was the standard normal distribution. Bollerslev (1987) tried to capture the high degree of leptokurtosis that is presented in high frequency data and proposed the Student-*t* distribution in order to produce an unconditional distribution with fat tails. Lambert and Laurent (2001) suggested that not only the conditional distribution of innovations may be leptokurtic, but also asymmetric and proposed the Skewed Student-*t* densities function.

According to Lambert and Laurent (2001) and provided that  $\upsilon > 2$ , the innovation process  $z_t$  is said to be (standardized) Skewed Student-*t* (in short SKST) distributed, i.e.  $z_t \square SKST(0,1,\xi,\upsilon)$ , if:

$$f(z_t|\xi,\upsilon) = \begin{cases} \frac{2}{\xi + \frac{1}{\xi}} sg[\xi(sz_t + m)|\upsilon] & \text{if } z_t < -\frac{m}{s} \\ \frac{2}{\xi + \frac{1}{\xi}} sg[(sz_t + m)/\xi|\upsilon] & \text{if } z_t \ge -\frac{m}{s} \end{cases}$$
(6)

where g(.|v) is a symmetric (unit variance) Student-*t* density and  $\xi$  is the asymmetric term. In short,  $\xi$  models the asymmetry, while v accounts for the tail thickness. Parameters *m* and  $s^2$  are, respectively the mean and the variance of the non-standardized Skewed Student-*t* density:

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$$m = \frac{\Gamma\left(\frac{\nu-1}{2}\right)\sqrt{\nu-2}}{\sqrt{\pi}\Gamma\left(\frac{\nu}{2}\right)} \left(\xi - \frac{1}{\xi}\right)$$
(7)

and

$$s^{2} = \left(\xi^{2} + \frac{1}{\xi^{2}} - 1\right) - m^{2}$$
(8)

Following Ding, Granger and Engle (1993), if it exists, a stationary solution of Equation (5) is given by:

$$E(\sigma_{t}^{\delta}) = \frac{\omega_{0}}{1 - \alpha_{1}E(|z| - \gamma z)^{\delta} - \beta_{1}}$$
(9)

which depends on the density of z. Such a solution exist if  $V = \alpha_1 E(|z| - \gamma z)^{\delta} + \beta_1 < 1$ . The V coefficient may be viewed as a measure of volatility persistence.

Ding, Granger and Engle (1993) derived the expression for  $E(|z| - \gamma z)^{\delta}$  for the Gaussian case. We can also show that for the standardized Skewed Student-*t* distribution is given as follows:

$$E(|z|-\gamma z)^{\delta} = \left\{ \xi^{-(1+\delta)} \left(1+\gamma\right)^{\delta} + \xi^{(1+\delta)} \left(1-\gamma\right)^{\delta} \right\} \frac{\Gamma\left(\frac{\delta+1}{2}\right) \Gamma\left(\frac{\upsilon-\delta}{2}\right) \left(\upsilon-2\right)^{\frac{1+\delta}{2}}}{\left(\xi+\frac{1}{\xi}\right) \sqrt{\left(\upsilon-2\right)\pi} \Gamma\left(\frac{\upsilon}{2}\right)}$$
(10)

It is possible to nest a number of the more standard ARCH and GARCH formulations within this Asymmetric Power GARCH model by specifying permissible values for  $\alpha, \beta, \gamma$  and  $\delta$  in Equation 4. Table 1 summarizes the restrictions required to produce each of the models nested within this APGARCH model. From Table 1, where  $\alpha_i$  is free,  $\delta = 2$  and both  $\beta_j$  and  $\gamma_i = 0$ , this model reduces to Engle's (1982) ARCH model. Further, when we extend this model to allow both  $\alpha_i$  and  $\beta_j$  to take on any value, we get Bollerslev's (1986) GARCH model. The GJR-ARCH model of Glosten, Jagannathan and Runkle (1993) is obtained where  $\delta = 2$  and  $\beta_j = 0$ . The Threshold ARCH (TARCH) model of Zakoian (1994) is defined where  $\alpha_i$  is free,  $\delta = 1$ ,  $|\gamma_i| \leq 1$  and  $\beta_j$  is restricted to be 0. The Nonlinear ARCH

model (NARCH) of Higgins and Bera (1992) is obtained where  $\delta$  and  $\alpha_i$  are free, and both  $\beta_i$  and  $\gamma_i$  are 0. If we extend this NARCH model to allow  $\beta_i$  to also being free, then a Power GARCH (PGARCH) specification is the result.

# Insert Table 1 about here

The models nested so far have assumed a symmetrical response of volatility to innovations in the market. However, empirical evidence suggests that positive and negative returns to the market of equal magnitude will not generate the same response in volatility. Glosten, Jagannathan and Runkle (1993) provided one of the first attempts to model asymmetric or leverage effects with a model which utilizes a GARCH type conditional variance specification. In this GJR-GARCH model,  $\delta = 2$  and  $\beta_j$  is free however,  $\alpha_i$  is specified as  $\alpha_i (1+\gamma_i)^2$  and leverage term is restricted to  $-4\alpha_i\gamma_i$ . The Generalized TARCH (TGARCH) model is derived by allowing  $\beta_j$  being free. Lastly, if  $\alpha_i$ ,  $\beta_j$  and  $\delta$  are free, and  $|\gamma_i| \le 1$ , then an Asymmetric Power GARCH specification is the result. Full details and proofs of this nesting process may be found in Ding, Granger and Engle (1993).

# DATA AND EMPIRICAL RESULTS

The section shows the empirical results of models. The closing prices of eight stock market price indices are analyzed. Computations were performed with G@RCH 4.2 which is Ox package designed for the estimation of various time series models. The characteristics of the data are presented in the first subsection. The second subsection shows the estimated results of APGARCH (1,1) Skewed Student-*t* model specifications and the corresponding qualification tests. To conserve space the GARCH (1,1) and GJR-GARCH (1,1) model results declined to present although they are available upon request. The APGARCH (1,1) model produced highly significant test statistics than GARCH (1,1) and GJR-GARCH (1,1) models. The SÜ İİBF Sosyal ve Ekonomik Araştırmalar Dergisi

APGARCH model contained either a significant asymmetry term or a power term which was significantly different from two.

# Data

The paper considers the national stock market closing prices for eight countries. These countries and their respective price indices are: USA (Nasdaq100), Germany (DAX), Japan (Nikkei225), Singapore (Strait Times), Argentina (MerVal), Mexico (IPC), China (Shanghai Composite) and Turkey (ISE100). The reason why these countries have been chosen is to reveal the disparity of results of the analysis and the applicability of the APGARCH (1,1) model in terms of developed (USA, Germany, Japan, Singapore) and emerging markets (Argentina, Mexico, China, Turkey). Country memberships for the Morgan Stanley Capital International (MSCI) International Equity Indices have taken into consideration for market distinction. The data obtained from the Yahoo Finance: World Indices database and the Istanbul Stock Exchange for the period 04.01.1999 to 02.07.2009. For each national stock market price indices, the continuously compounded rate of return was estimated as  $r_t = \ln(p_t / p_{t-1})$  where  $p_t$  is the closing price on day t.

The usual descriptive statistics for each stock market return series are summarized in Table 2. It is not surprising that these series exhibit asymmetric and leptokurtic (fat tails) properties. Thus, the return series of these stock indices are not gauss distributed. Also the Jarque-Bera statistic is highly significant for each of the models indicating non-normality of the data. The JAP, SNG, ARG and CHN stock market returns are negatively skewed while the others positively.

# Insert Table 2 about here

From the descriptive graphics presented in Figure 1, several volatility periods can be observed. These graphical expositions show that all of the return series exhibit volatility clustering which means that there are periods of large absolute changes tend to cluster together followed by periods of relatively small absolute changes.

# **Insert Figure 1 about here**

### **Estimation Results**

In this subsection, the APGARCH (1,1) model is estimated for each national stock market return series under Gauss, Student-*t*, GED (Generalized Error Distribution) and Skewed Student-*t* distributions. The standard of model selection is based on in-sample diagnosis including Akaike Information Criterion (AIC), log-likelihood (LL) values, and Ljung-Box Q and  $Q^2$  statistics on standardized and squared standardized residuals respectively. Under every distribution, the model which has the lowest AIC or highest LL values and passes the Q-test simultaneously is adopted. In summary, ranking by AIC and LL favours the APGARCH (1,1) Skewed Student-*t* specification in all national stock market return series.

Table 3 presents the results of this estimation procedure and from this table one can see that all of the ARCH and GARCH coefficients are statistically significant at the 1% level. Further, the sum of the ARCH and GARCH coefficients for all of the models estimated was less than unity indicating that shocks to the model are transitory rather than permanent. Also,  $\beta_1$  is close to 1 but significantly different from 1 for all series, which indicates a high degree of volatility persistence.  $\beta_1$  takes values between 0.87 to 0.95 suggesting that there are substantial memory effects. Furthermore, in USA and GER cases the APGARCH models are persistent in the sense that *V* coefficient is equal to 1, and in all other cases the APGARCH models are stationary in the sense that *V* coefficient is lower than 1.

The APGARCH model includes a leverage term ( $\gamma$ ) which allows positive and negative shocks of equal magnitude to elicit an unequal response from the market. Table 3 presents details of this leverage term and reveals that for all models fitted; the estimated coefficient was positive and statistically significant. This means that negative shocks lead to higher subsequent volatility than positive shocks (asymmetry in the conditional variance). Such a result was expected since response asymmetry is generally attributed solely to stock market data. From Table 3, the evidence of long memory process could be also found in the results of the model estimation because the power term ( $\delta$ ) of APGARCH models range in value from 1.9044 in the case of ARG to 1.2948 in the case of JAP. The average power term across all of the models estimated was 1.5796. For two of the models (MEX and JAP) estimated the power term was significantly different from two and for an additional four models (USA, SNG, ARG and TUR) estimated the power term was significantly different from unity. This means that for six of the eight models estimated, the optimal power term was some value other than unity or two which would seem to support the use of a model which allows the power term to be estimated. The APGARCH models the conditional standard deviation for the MEX and JAP return series and the conditional variance for the USA, SNG, ARG and TUR return series.

For the Skewed Student-*t* distribution, the asymmetric terms are negative ( $\xi$ <0) and statistically significant for all the national stock market return series except TUR. Note that G@RCH does not estimate  $\xi$  but log( $\xi$ ) to facilitate inference about the null hypothesis of symmetry (since the Skewed Student-*t* equals the symmetric Student-*t* distribution when  $\xi$ =1 or log( $\xi$ )=0). The sign of log ( $\xi$ ) indicates the direction of the skewness. The third moment is positive, and the density is skew to the right, if log ( $\xi$ )>0. On the contrary, the third moment is negative, and the density is skew to the left, if log ( $\xi$ )<0. We can confirm that the density distributions of all series are skewed to the left side due to these significantly negative asymmetric terms.

The tail term (v) is much larger for the USA, GER and JAP returns than for the other series. This means that daily returns of the SNG, ARG, MEX, CHN and TUR stock market price indices display a much larger kurtosis and exhibit fatter tails than returns for the USA, GER and JAP stock market price indices. Besides, the evidences show that fat-tail phenomenon is strong because the student or tail terms (v) are significantly different from zero for all series under Skewed Student-*t* distribution.

#### Insert Table 3 about here

The results given in Table 3 show that the APGARCH succeeds in taking into account all the dynamical structure exhibited by the returns and volatility of the returns as the Ljung-Box statistics for up to 20 lags on the standardized residuals (Q) non-significant at the 5% level (except USA, MEX and CHN return series) and the squared standardized residuals ( $Q^2$ ) non-significant at the 5% level (except GER and JAP return series).

#### CONCLUSION

A recent development in the ARCH literature has been the introduction of the Power ARCH class of models which allow a free power term rather than assuming an absolute or squared term in their specification. Accordingly, the purpose of the current paper was to consider the applicability of the Generalized Asymmetric Power ARCH (APGARCH) model to the selected national stock market returns for eight countries. The stock market indices were investigated by using the APGARCH (1,1) model with Skewed Student-*t* distribution. To capture the long memory property exhibited in the conditional variance, the power term ( $\delta$ ) estimates of APGARCH model is in the interval between one and two. It indicates that the return series of all the national stock market price indices are skewed distributed and have fat tails by the significant coefficients of  $\xi$  (not significant for TUR) and v in the results of model estimation. The skewed Student-*t* density appears to be a promising specification to accommodate both the high kurtosis and the skewness inherent to most asset returns.

The estimation results indicate that strong leverage effects are present in national stock market data especially for USA, GER, JAP and MEX. Further, once these leverage effect are modeled in a GARCH framework, the inclusion of a power term is a worthwhile addition to the specification of the model. Also, in developed markets the volatility persistence is higher than emerging markets. Thus, shocks in the return series have substantial memory effects.

Consequently, in this paper, the ability of the APGARCH model is analyzed so as to present the volatility characteristics of four developed and four emerging markets. The results suggest that in the presence of asymmetric responses to innovations in the market, the APGARCH(1,1) Skewed Student-*t* model which accommodates both the skewness and the kurtosis of financial time series is favored. However, as internal dynamics of each market are different, it is concluded that there is no possibility to judge and to interpret the parameters of the model by differentiating markets as developed and emerging markets.

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MODEL	$lpha_i$	$\beta_{j}$	γi	δ
ARCH	Free	0	0	2
GARCH	Free	Free	0	2
GJR-ARCH	$\alpha_i(1{+}\gamma_i)^2$	0	$-4\alpha_i\gamma_i$	2
GJR-GARCH	$\alpha_i(1+\gamma_i)^2$	Free	$-4\alpha_i\gamma_i$	2
TARCH	Free	0	$\left   \gamma_i  \right   \leq 1$	1
TGARCH	Free	Free	$\left   \gamma_i  \right   \leq 1$	1
NARCH	Free	0	0	Free
PGARCH	Free	Free	0	Free
APARCH	Free	0	$\left   \gamma_i  \right   \leq 1$	Free
APGARCH	Free	Free	$\left   \gamma_i  \right   \leq 1$	Free

Table 1. Taxonomy of ARCH/GARCH model specifications

**Table 2. Descriptive statistics** 

	USA	GER	JAP	SNG	ARG	MEX	CHN	TUR
Observation	2,640	2,668	2,577	2,577	2,595	2,631	2,456	2,613
Minimum	-0.1112	-0.0743	-0.1211	-0.0922	-0.1295	-0.0827	-0.0926	-0.1998
Maximum	0.1720	0.1080	0.1324	0.0753	0.1612	0.1044	0.0940	0.1777
Standard Deviation	0.0225	0.0166	0.0162	0.0137	0.0227	0.0159	0.0166	0.0270
Skewness	0.2095	0.0634	-0.2918	-0.3839	-0.0323	0.1424	-0.0229	0.1115
Excess Kurtosis	3.9101	4.0828	6.2717	5.1401	4.8640	3.4524	4.6393	4.9709
Jarque-Bera	1,701	1,855	4,260	2,900	2,559	1,316	2,203	2,696
(Prob.)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)
ADF-Test (C, 0)*	-54.92	-53.31	-52.71	-49.43	-48.32	-45.99	-49.28	-50.16

**Note:** USA: Nasdaq100 (USA), GER: DAX (Germany), JAP: Nikkei225 (Japan), SNG: Strait Times (Singapore), ARG: MerVal (Argentina), MEX: IPC (Mexico), CHN: Shanghai Composite (China), TUR: ISE100 (Turkey).

\* (C, 0) indicates that there is a constant but no trend in the regression model with lag=0. All Augmented Dickey Fuller (ADF) test statistics reject the hypothesis of a Unit Root at the 1% level of confidence. MacKinnon critical value at the 1% confidence level is -3.44.



**Note:** USA: Nasdaq100 (USA), GER: DAX (Germany), JAP: Nikkei225 (Japan), SNG: Strait Times (Singapore), ARG: MerVal (Argentina), MEX: IPC (Mexico), CHN: Shanghai Composite (China), TUR: ISE100 (Turkey).



	USA	GER	JAP	SNG	ARG	MEX	CHN	TUR
μ	0.0002	0.0002	0.0000	0.0004	0.0006	0.0008	0.0002	0.0011
	[0.619]	[0.671]	[0.167]	[2.091 <sup>a</sup> ]	[1.563]	[3.558]	[0.670]	[2.759]
ω	0.0253	0.2428	0.8868	0.1120	0.1911	0.7723	0.3456	0.3122
	[0.767]	[0.903]	[0.888]	[0.734]	[0.756]	[0.850]	[0.914]	[0.714]
α	0.0438	0.0659	0.0833	0.0936	0.1039	0.0835	0.1049	0.1062
	[4.008]	[5.792]	[7.809]	[5.498]	[6.126]	[6.581]	[3.670]	[5.057]
β	0.9474	0.9187	0.9086	0.9078	0.8715	0.9070	0.9085	0.8843
	[120.3]	[99.1]	[79.2]	[56.3]	[41.1]	[63.9]	[35.0]	[38.86]
Ŷ	0.5590	0.7083	0.4795	0.2214	0.2293	0.5665	0.2065	0.1650
	[3.183]	[3.856]	[3.848]	[3.085]	[3.474]	[5.202]	[3.162]	[2.774]
δ	1.7933	1.4978	1.2948	1.5887	1.9044	1.3315	1.4662	1.7603
	[7.167]	[6.611]	[5. 495]	[5.816]	[5.568]	[5.290]	[6.808]	[5.014]
ξ	-0.0654	-0.1231	-0.0671	-0.0446	-0.0535	-0.0491	-0.0449	0.0139
	[-2.468 <sup>a</sup> ]	[-4.653]	[-2.566 <sup>a</sup> ]	[-1.632 <sup>b</sup> ]	[-2.074 <sup>a</sup> ]	[-1.894 <sup>b</sup> ]	[-1.993 <sup>a</sup> ]	[0.473 <sup>c</sup> ]
υ	20.7822	14.7119	13.0248	7.1105	5.2517	8.3216	3.8336	7.4105
	[2.560 <sup>a</sup> ]	[3.276]	[3.925]	[7.208]	[9.813]	[6.615]	[10.820]	[7.576]
V	1.0052	1.0059	0.9861	0.9929	0.9828	0.9849	0.9929	0.9811
LL	6,925.16	7,733.81	7,378.72	7,873.23	6,596.87	7,552.66	7,028.84	6,117.44
AIC	-5.2403	-5.7915	-5.7204	-6.1042	-5.0781	-5.7352	-5.7173	-4.6762
Q(20)	32.6182	27.0778	18.4249	18.2216	30.8445	48.1371	39.0405	29.5567
	(0.0371)	(0.133)	(0.559)	(0.573)	(0.0573)	(0.000)	(0.007)	(0.077)
$O^2(20)$	28.5914	71.6271	32.4275	17.7468	16.3753	17.0795	8.7155	20.6688
Q*(20)	(0.0536)	(0.000)	(0.020)	(0.472)	(0.566)	(0.518)	(0.966)	(0.296)

Table 3. APGARCH (1,1) skewed Student-t model estimation results

**Notes:** USA: Nasdaq100 (USA), GER: DAX (Germany), JAP: Nikkei225 (Japan), SNG: Strait Times (Singapore), ARG: MerVal (Argentina), MEX: IPC (Mexico), CHN: Shanghai Composite (China), TUR: ISE100 (Turkey).

a, b denote 5% and 10% significance level respectively; c denotes insignificancy;  $V = \alpha_i E(|z| - \gamma z)^{\delta} + \beta_i$  as a measure of volatility persistence, t-statistics of corresponding tests in brackets. AIC-Akaike Information Criterion, LL is the value of the maximized log-likelihood. Q(20) and Q<sup>2</sup>(20) are the Ljung-Box statistics for remaining serial correlation in the standardized and squared standardized residuals respectively using 20 lags with p-values in parenthesis.